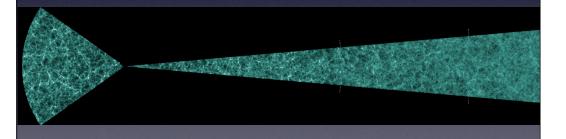


halo model paradigm of cosmological large-scale structure (LSS)

LSS = a hierarchical web of quasi-equilibrium bound structures - *halos* - that emerge via gravitational amplification from a noise field imposed during an early epoch of inflation.

Halo Model's key enabling ingredients:

- space density (aka, mass function), n(M, z)
- spatial N-point correlations (e.g., 2-pt bias function), b(M, z)
- internal halo structure (kinematics, thermodynamics), $X(r/r_{\Delta}, M, z)$



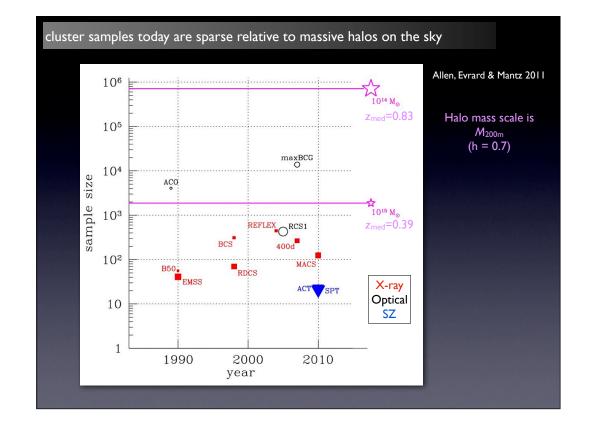
the lexicon

* Halo:

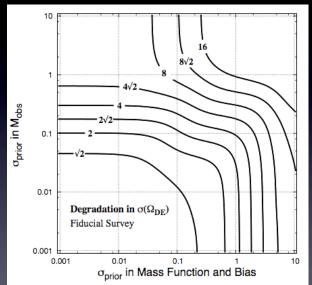
a self-bound, quasi-equilibrium structure comprised of multiple, interacting fluids (dark matter, multi-phase baryons, and radiation) formed via gravitational collapse within a cosmic web of random noise.

* Cluster:

a redshift-space projection of a massive halo, and its line-of-sight neighbors, with the resultant system containing multiple, bright galaxies and other visible components (multi-phase baryons, non-thermal matter, etc.).



Fisher matrix motivation to understand scaling relations

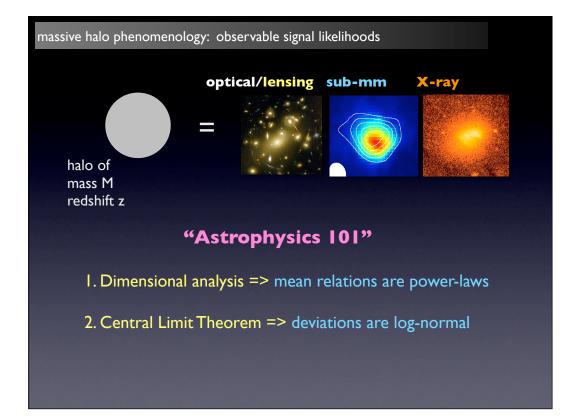


Cunha + AE (2009)

scaling relation (Mobs) uncertainty is presently the dominant source of systematic error in DE studies

a PL+LN* multivariate signal model

*Power Law + Log-Normal



1 A Local Model for Multivariate Counts

Consider a mass function described locally as a power-law in mass with slope $-\alpha$. Specifically, using $\mu \equiv \ln M$, define the mass function, $n(\mu, z)$, as the likelihood of finding a halo at redshift z in the mass range μ to $\mu + d\mu$ within a small comoving volume dV.

$$dp \equiv n(M, z) \, d\ln M \, dV = A M^{-\alpha} d\ln M \, dV = A e^{-\alpha \mu} \, d\mu \, dV. \tag{1}$$

The local slope, α , and amplitude, A, implicitly depend on mass and redshift in a manner dependent on cosmology (e.g., Tinker et al. 2008).

Consider a set of N halo properties, $S_i \in \{N_{\text{gal}}, L_{\text{X}}, T_{\text{X}}, M_{\text{gas}}, Y_{\text{X}}, Y_{\text{SZ}}, \cdots\}$, let \mathbf{s} be a vector containing their logarithms,

$$s_i = \ln(S_i) \tag{2}$$

Assume that the mass scaling behavior of these properties are power-laws, so that the mean ln(signal) for a mass–complete sample scales as

$$\bar{\mathbf{s}}(\mu, z) = \mathbf{m}\mu + \mathbf{b}(z). \tag{3}$$

The elements of vector **m** are the slopes of the individual mass-observable relations. (Note that, at some fixed epoch, we can always choose units such that the intercepts $b_i(z) = 0$.)

Assume that ln(signal) deviations about the mean are Gaussian, described by a likelihood

$$p(\mathbf{s}|\mu) = \frac{1}{(2\pi)^{N/2} |\Psi|^{1/2}} \exp\left[-\frac{1}{2} (\mathbf{s} - \overline{\mathbf{s}})^{\dagger} \Psi^{-1} (\mathbf{s} - \overline{\mathbf{s}})\right], \tag{4}$$

where the covariance matrix has elements

$$\Psi_{ij} \equiv \langle (s_i - \bar{s}_i)(s_j - \bar{s}_j) \rangle, \tag{5}$$

and the brackets denote an ensemble average over a (large) mass-complete sample.

1.1 Multivariate Space Density

The space density as a function of the multivariate properties, s, is found by the convolution, $n(\mathbf{s}) = \int d\mu \, n(\mu) \, p(\mathbf{s}|\mu)$. Using equations (1) and (4), the result is

$$n(\mathbf{s}) = \frac{A\Sigma}{(2\pi)^{(N-1)/2} |\Psi|^{1/2}} \exp\left[-\frac{1}{2} (\mathbf{s}^{\dagger} \Psi^{-1} \mathbf{s} - \frac{\bar{\mu}^2(\mathbf{s})}{\Sigma^2})\right],\tag{6}$$

where Σ^2 is the multi-property mass variance defined by

$$\Sigma^2 = (\mathbf{m}^{\dagger} \Psi^{-1} \mathbf{m})^{-1}, \tag{7}$$

and the mean mass is

$$\bar{\mu}(\mathbf{s}) = \frac{\mathbf{m}^{\dagger} \Psi^{-1} \mathbf{s}}{\mathbf{m}^{\dagger} \Psi^{-1} \mathbf{m}} - \alpha \Sigma^{2},$$

$$\equiv \bar{\mu}_{0}(\mathbf{s}) - \alpha \Sigma^{2}.$$
(8)

$$\equiv \bar{\mu}_0(\mathbf{s}) - \alpha \Sigma^2. \tag{9}$$

The first term, $\bar{\mu}_0(\mathbf{s})$, is the mean mass for the case of a flat mass function, $\alpha = 0$, which corresponds to the mass expected from inverting the input log-mean relation.

The second term, $\alpha \Sigma^2$, represents the mass shift induced by asymmetry in the convolution when $\alpha > 0$. (Low mass halos scattering up outnumber high mass systems scattering down.) Note that the magnitude of this effect scales with the variance, not the rms deviation.

Applying Bayes' theorem in the form $p(\mu|\mathbf{s}) = p(\mathbf{s}|\mu)n(\mu)/n(\mathbf{s})$ leads to the result that the set of masses selected by a specific set of properties is Gaussian in the log with mean given by equation (9) and variance, equation (7).

1.1.1 Explicit expressions for the one-variable case

For a single property, $s \equiv \ln(S)$, with slope, m, and logarithmic scatter at fixed mass, σ , the mass variance at fixed S is

$$\Sigma^2 = \left(\frac{\sigma}{m}\right)^2. \tag{10}$$

The mean mass for a sample complete in S is

$$\bar{\mu}(s) = \frac{s}{m} - \alpha \Sigma^2. \tag{11}$$

The property space density function is

$$n(s) ds = (A/m) \exp\{-\alpha \left(\frac{s}{m} - \alpha \Sigma^2 / 2\right)\} ds, \tag{12}$$

which is a power-law in the original property, $n(S) \propto S^{-(\alpha/m)}$.

Note that the effective shift in mass, $\alpha \Sigma^2/2$, is half that in the expression above. These expressions are consistent, in that they address different questions. Equation (11) gives the mean ln(mass) of a signal-selected sample while equation (12) gives the ln(mass) value that matches the local space density – in number per volume per ln(S) – of halos with property value, S.

1.1.2 Explicit expressions for the two-variable case

For two properties, we introduce the correlation coefficient, $r \equiv \langle \delta_1 \delta_2 \rangle$, of the normalized deviations, $\delta_i \equiv (s_i - \bar{s}_i)/\sigma_i$, and write the covariance matrix,

$$\Psi = \left(egin{array}{cc} \sigma_1^2 & r\sigma_1\sigma_2 \ r\sigma_1\sigma_2 & \sigma_2^2 \end{array}
ight),$$

and its inverse,

$$\Psi^{-1} = (1 - r^2)^{-1} \begin{pmatrix} \frac{1}{\sigma_1^2} & -\frac{r}{\sigma_1 \sigma_2} \\ -\frac{r}{\sigma_1 \sigma_2} & \frac{1}{\sigma_2^2} \end{pmatrix}.$$

The mass variance is now a harmonic mixture

$$\Sigma^{-2} = (1 - r^2)^{-1} \left(\sigma_{u1}^{-2} + \sigma_{u2}^{-2} - 2r\sigma_{u1}^{-1}\sigma_{u2}^{-1} \right), \tag{13}$$

where $\sigma_{\mu i} = \sigma_i/m_i$ is the mass scatter at fixed signal S_i .

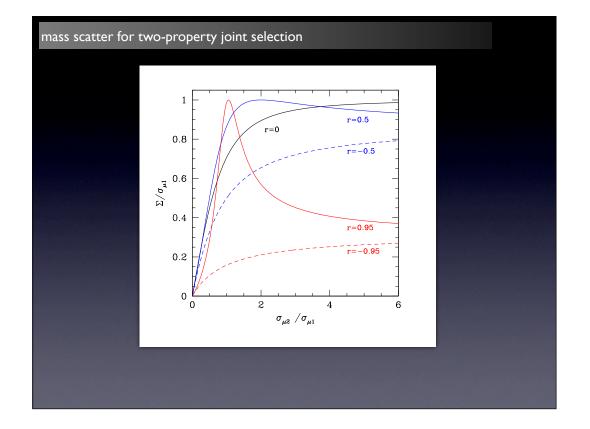
The zero-slope mean mass is

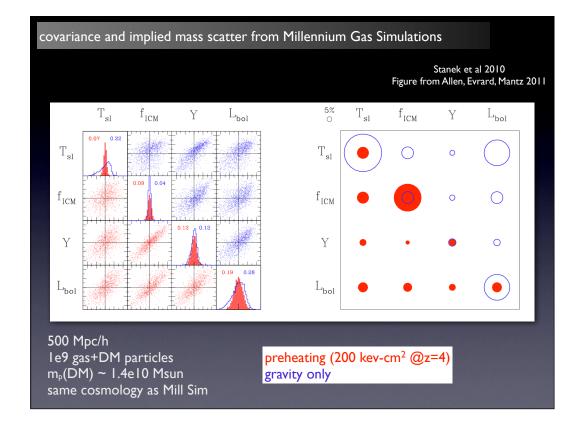
$$\bar{\mu}_0(s_1, s_2) = \frac{(s_1/m_1)\sigma_{\mu 1}^{-2} + (s_2/m_2)\sigma_{\mu 2}^{-2} - r(s_1/m_1 + s_2/m_2)\sigma_{\mu 1}^{-1}\sigma_{\mu 2}^{-1}}{\sigma_{\mu 1}^{-2} + \sigma_{\mu 2}^{-2} - 2r\sigma_{\mu 1}^{-1}\sigma_{\mu 2}^{-1}},$$
(14)

and the joint space density is

$$n(s_1, s_2) = \frac{A\Sigma}{\sqrt{2\pi(1 - r^2)}\sigma_1\sigma_2} \exp\left[-\alpha\bar{\mu}_0 + \frac{\Sigma^2}{2}\left(\alpha^2 - \frac{(s_1/m_1 - s_2/m_2)^2}{\sigma_{\mu 1}^2\sigma_{\mu 2}^2}\right)\right].$$
(15)

The first two terms in the exponent are analogous to those in the 1D expression, equation (12). For "reasonable" choices of (S_1,S_2) pairs — meaning values that pick out comparable mass scales, $s_1/m_1 \sim s_2/m_2$ — the space density remains effectively power-law. The third term in the exponent suppresses the number density for unreasonable pairings of s_1/m_1 and s_2/m_2 , those lying out in the wings of the bivariate Gaussian.





1.2 Property-selected samples

For a halo sample selected with some property, s_1 , we can now use Bayes' theorem to find the joint probability of those halos having a second property, s_2 , and mass, μ . The result can be expressed as a bivariate Gaussian in terms of the two-element vector, $\mathbf{t} = [s_2 \ \mu]$,

$$p(\mathbf{t}|s_1) = \frac{1}{(2\pi)|\tilde{\Psi}|^{1/2}} \exp\left[-\frac{1}{2}(\mathbf{t}-\bar{\mathbf{t}})^{\dagger}\tilde{\Psi}^{-1}(\mathbf{t}-\bar{\mathbf{t}})\right], \tag{16}$$

where the mean mass, $\bar{\mu}(s_1)$, is defined by equation (11) and the mean of the non-selection property is given by

$$\bar{s}_2(s_1) = m_2 \left(\bar{\mu}(s_1) + \alpha r \sigma_{\mu 1} \sigma_{\mu 2} \right).$$
 (17)

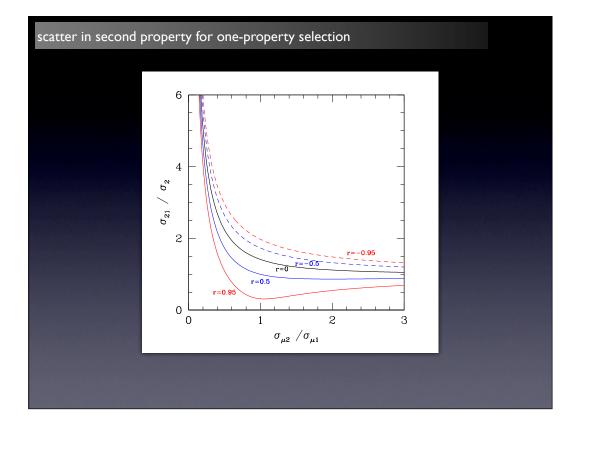
Note that, if r < 0, the non-selected property mean can be "doubly" biased low relative to a simple $m_2(s_1/m_1)$ expectation, with one shift coming from the extra $(-\alpha \Sigma^2)$ term in the mean mass and the second coming from the second term in the above expression.

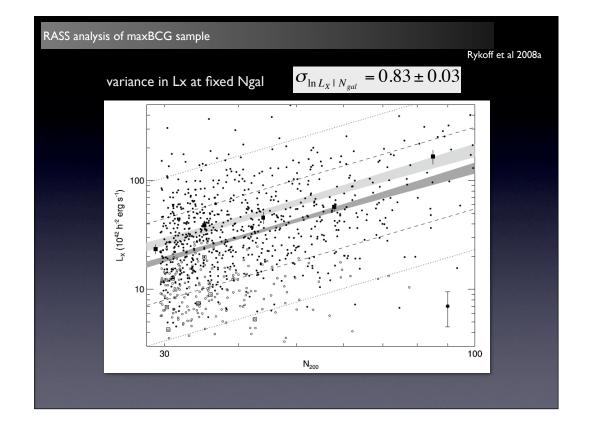
The covariance in s_2 and μ at fixed s_1 is given by

$$ilde{\Psi} = \left(egin{array}{cc} \sigma_{21}^2 & ilde{r}\sigma_{21}\sigma_{\mu 2} \ ilde{r}\sigma_{21}\sigma_{\mu 2} & \sigma_{\mu 2}^2 \end{array}
ight),$$

where the variance in s_2 at fixed s_1 is

$$\sigma_{21}^2 = m_2^2 \left(\sigma_{\mu 1}^2 + \sigma_{\mu 2}^2 - 2r \sigma_{\mu 1} \sigma_{\mu 2} \right). \tag{18}$$



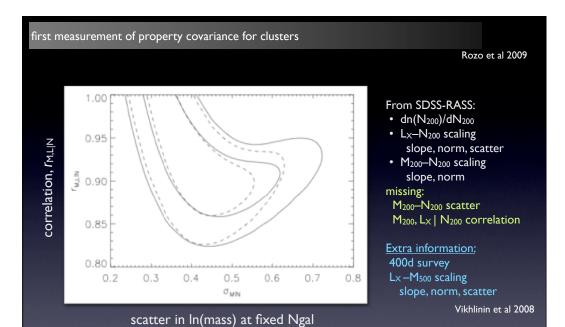


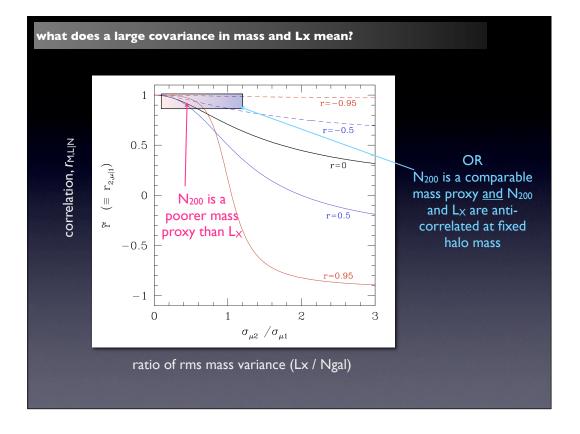
The s_2 -mass correlation coefficient, \tilde{r} , depends on both the intrinsic property correlation, r, as well as the ratio of scatter in mass for the two properties,

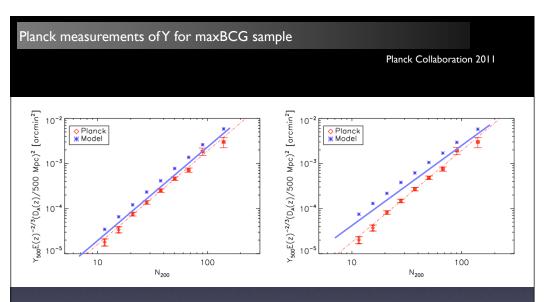
$$\tilde{r} = \frac{\sigma_{\mu 1}/\sigma_{\mu 2} - r}{\sqrt{1 - r^2 + (\sigma_{\mu 1}/\sigma_{\mu 2} - r)^2}}.$$
(19)

If the selection property is an excellent mass proxy $(\sigma_{\mu 1} \to 0)$, then $\tilde{r} \to -r$.

If the selection property is a much poorer mass proxy compared to the second property, then $\tilde{r} \to 1$, irrespective of the intrinsic correlation, r.







One likely error source: model assumes $M_{X-ray} = M_{lens}$

Blue lines: $M_{X-ray} = 0.8 M_{lens}$

summary

era of large, overlapping multi-wavelength surveys

- is getting nearer
- will enable stringent tests of the basic PL+LN model Are we doing all we can now?

note that

- selection effects (that include projection) must be carefully modeled
- effect of projection on covariance measurements needs study

How to "world average" slopes, intercepts, covariance? What are the limits of the basic PL + LN model?